NOWCASTING GDP: A note on forecasting improvements from the use of bridge models Evidence from Greece and six other European Countries

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Abstract

In the recent literature on nowcasting, the use of the so-called bridge models is advocated. These are simple regression models that use data on mixed frequencies, usually with the lower frequency data serving as dependent variables and the higher frequency data as explanatory variables. In this note we investigate whether the use of such models can lead to performance enhancements in forecasting real GDP growth for Greece and six other European countries. This is an interesting and instructive exercise because of the obvious break in Greek real GDP growth during the crisis but also, and more importantly, because of the potential usefulness of such models in forecasting the anticipated recovery in Greek growth. Applying the exercise in the GDP growth rates of Cyprus, Ireland, Germany, Portugal, Spain and UK we confirmed our findings. Since many monthly activity indicators are released in advance of GDP growth it is interesting to see how the structure and timing of bridge models can lead to potential improvements in forecasting growth. Our results indicate that by using three of the most important monthly activity indicators such performance enhancements are indeed possible.

KEYWORDS: bridge models, nowcasting, GDP, Greece, growth

JEL classification: C52, C53, E01, E27

1. Introduction

Decision-makers in different parts of the economy such as business, government, the central bank, financial markets and others, are in need of an accurate and timely assessment of economic growth. The main problem is that since most macroeconomic series of interest are only available at a quarterly frequency and are released three to six weeks after the close of the quarter, many institutions are faced with the problem of using monthly information in order to obtain an early estimate of the last quarter and the current quarter results, as well as a forecast for one quarter ahead.

The aim of this paper is to attempt a nowcasting exercise for the Greek real growth rate by exploiting the particular structure of data on the Greek economy and their release. What makes our exercise particularly interesting is the problems of the data themselves and the importance of growth assessments and forecasts in the context of the deep fiscal crisis faced by the Greek government and productive sectors.

Moreover, we applied our exercise in six other European countries: Cyprus, Ireland, Germany, Portugal, Spain and UK in order to confirm that our method is valid not only for the peculiarity of the Greek data but for the Growth Rates of the other countries.

Nowcasting is a relatively new method whose main advantage is the use of new information as it comes in, and the generation of updates at a higher frequency than the frequency of observation of the variable of interest. Until recently, nowcasting had received very little attention in the academic literature, although it was routinely conducted in policy institutions either through a judgmental process or on the basis of simple models. It was first introduced by Evans (2005) for a limited number of time series and evolved by Giannone, Reichlin, and Small (2008) for a larger number of

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series. In recent years, there have been many applications of this method for several countries and variables thus enhancing and expanding this methodology, such as Antonello *et al.* (2008) for Ireland.

In order to have better forecasts, factor models have proved to be a very useful tool for short-term forecasting of real activity. The use of dynamic factor models has been further improved by recent advances in estimation techniques proposed by Stock and Watson (2002a, 2002b), Forni *et al.* (2004, 2005) or Giannone, Reichlin, and Small (2008), who have put forward the advances in estimation techniques that allow improving their efficiency. This type of model is particularly appealing as it can be applied to large data sets [e.g., Angelini, Camba-Mendez, Giannone, Reichlin, & Rünstler (2011); Barhoumi, Darné, & Ferrara (2010); Schumacher & Breitung (2008); Schumacher (2007)].

The DFMs are based on static and dynamic principal components. The static principal components are obtained as in Stock and Watson (2002a, 2002b). The dynamic principal components are based on either time domain methods, as in Doz, Giannone and Reichlin (2011, 2012), or frequency domain methods, as in Forni *et al.* (2004, 2005). To the best of our knowledge, Banerjee, Marcellino, and Masten (2005), Banerjee and Marcellino (2006), Antipa *et al.* (2012) are the only studies that compare the forecasting performance of the automatically selected BMs and the DFMs – for Eurozone, US and German GDP growth, respectively. These studies, however, only use factor models following Stock and Watson (2002a, 2002b), for which results are not conclusive in favor of one or the other. DFMs have so far never been used for forecasting Greek GDP growth rates. While the econometric performance of DFMs is very satisfactory, an important caveat of this approach is that the economic content of factors is difficult to interpret from an economic point of

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view. For that reason we complete this analysis by several bridge models which allow for a more straightforward interpretation of the data used.

An alternative approach to the analysis of time series with mixed frequencies is the mixed data sampling regression (MIDAS) method proposed by Ghysels, Santa-Clara, and Valkanov (2006). The MIDAS method provides linear projections without specifying the dynamics of the regressors. When the model is specified correctly and the parameters are known, the Kalman filter is superior to MIDAS by construction. Otherwise, the question of whether MIDAS or the state space method is superior is still under investigation; see the study of Bai, Ghysels, and Wright (2011), who consider both MIDAS and state space methods. They show the conditions under which the methods are identical and provide evidence that the Kalman filter is slightly more accurate.

The rest of the paper is organized as follows. In section 2 we give a brief summary of the bridge models. In section 3 we discuss the results of our forecasting analysis and section 4 offers some concluding remarks for future research.

2. The bridge model & data, estimation and forecasting

Bridge models are essentially mixed frequency linear regressions. These models "bridge", i.e. link, monthly variables to quarterly ones - hence their name. In this sense they are unrestricted versions of the MIDAS approach (Ghysels, Santa-Clara, and Valkanov (2006)). Such models have been widely considered in the recent literature, and are especially used to forecast GDP growth in national and international institutions (e.g. Diron, 2008; Golinelli & Parigi, 2005; Parigi &

Schlitzer, 1995; Rünstler & Sédillot, 2003; Sédillot & Pain, 2003; Zheng & Rossiter, 2006).

To make things specific, let us consider monthly and quarterly variables in the context of our data. The explanatory variables will be monthly economic activity indicators, namely the index of industrial production (IPI), the total turnover of retail sales (RSTOT) and the volume of retail sales (RSVOL). All variables are from seasonally adjusted indices and expressed in real terms as annual growth rates. The dependent variable is obtained from the, seasonally adjusted, quarterly real GDP series and also expressed as annual growth rate. All variables for the Greek economy are obtained from the Greek Statistical Authority website (<u>www.statistics.gr</u>) and for the rest six countries from the European Commission's statistical authority –EUROSTAT-(<u>http://ec.europa.eu/eurostat</u>)

-Table 1-

Data availability is from 2001 for real GDP and this dictates the rest of our analysis: we split the data into a training period up to 2007 and use the post-crisis data as our evaluation period.

The Real GDP Growth rate for Greece varies from -0,0894 at the third quarter of 2010, which is the trough, to the peak 0,0754 at the second quarter of 2006. The variable which is most correlated with the GDP is the Volume of Retail Sales of the previous month of examination, followed by the Total Turnover of Retail Sales of the previous month of examination. As can be seen in Table 2 there is a negative skewness between the variables and the values are wider spread around the mean.

-Table 2-

The estimation is conducted recursively to fully utilize the relatively small amount of observations available.

The general specification of a bridge model is that of an autoregressive-distributed-lag (ARDL) for q explanatory variables and is given as follows:

$$Y_{t} = a + \sum_{i=1}^{m} \beta_{i} Y_{t-i} + \sum_{j=1}^{q} \sum_{i=1}^{k} \delta_{j,i} X_{j,t-1} + \varepsilon_{t}$$

where *m* is the number of autoregressive parameters, *q* is the number of explanatory variables, and *k* is the number of lags for the explanatory variables. Note that under the restriction that now monthly variables appear above, we see that the equation collapses to a standard autoregression – which thus becomes the natural benchmark to compare forecasting performance. In our analysis we consider models that use each monthly variable, a pair of monthly variables and all three monthly variables together. These models are benchmarked against an AR(1) model and an AR(AIC) model, with maximum lags set to 6.

An important point we should make is that we use our data aligned correctly and taking account of release lags. This is important for making the exercise realistic. For example, we always use a two-month lag on the aligned monthly data: if we are at the end of the 4th quarter we use monthly data for October. So, if the real GDP for the 4th quarter is released, for example, in mid-February and the monthly variable is released in November or December we always use past data correctly in producing the forecasts.

	Qua	arter t			Quarter t+1					
							Real	GDP		
							release	ed here		
October	Nove	mber	Dece	mber	Janu	uary	Feb	ruary	Ma	rch
Monthly	Dei			ind						
data used	Data collection period									

Finally, to evaluate our forecasting results we use the standard measures of mean forecasting error, mean squared error and mean absolute error.

3. Forecasting results

3.1 Forecasting Results for Greece

Results in terms of mean error (ME), mean absolute error (MAE), mean-squared error (MSE) and root mean-squared error (RMSE) of the forecasts, as presented in Table 1 as well as the ratio obtained from AR(1) (Ratio1) and AR(AIC) (Ratio 2) benchmarks show that the combination of the IPI, the RSVOL and the RSTOT performed better than the benchmarks. Both Ratio1 and Ratio 2 showed that almost all models - except for the IPI- perform better than the benchmarks.

-Table 3-

Obviously, simply comparing error-values does not take into account the sample uncertainty underlying observed forecast differences. To ascertain if the differences in predictive accuracy found are statistically significant, we conduct formal testsbased on the Diebold and Mariano (1995) test statistic. In particular, we employ the following small sample version (DM) proposed by Harvey, Leybourne, and Newbold (1997):

$$DM = \sqrt{\frac{T + 1 - 2h + \frac{h(h - 1)}{T}}{T}} \frac{d}{\sqrt{V(d)}}$$
$$\sqrt{V(d)} = \frac{1}{T} (\gamma_0 + 2\sum_{k+1}^{h-1} \gamma_k)$$

where T is the number of forecasts made, h is the forecast horizon in months, d is the mean difference between the squared (or alternatively, absolute) forecast errors from any two competing models, V(d) is the approximate asymptotic variance of d , and applied the test of equality of forecast performance proposed by Diebold and Mariano (1995).

-Table 4-

Table 4 includes the results of Diebold–Mariano tests for equality of mean squared errors of each pair of forecasts for each individual series for the reported horizons. As can be seen the results are not as accurate as we would have expected, owing to the small amount of observations. The combination of the three models appears to have the best results over the AR(1) model.

3.2 Forecasting Results for Cyprus, Ireland, Germany, Portugal, Spain and UK

In order to prove that the method used, in the case of Greece, can be generally applied in other countries, we used a sample of 6 European countries: Cyprus, Ireland, Germany, Portugal, Spain and UK.

-Table 5-

Due to lack of data, in the case of Portugal, Cyprus and Ireland we used only two out of the three variables, the Volume of the sales and Industrial Production Index. The RSVOL for Portugal. with RMSE at 0.012, seems to perform better over the benchmark as well as the IPI and the combination of both variables. In the case of Cyprus, we can see that both the RSVOL and the combination of the variables perform better than the benchmark. For Ireland as we can see by the ratio 2 which is the ratio between the RMSE of each variable over the Benchmark AR(AIC) only RSVOL and IPI are greater than 1.

-Table 6-

To evaluate our results we used the Mincer and Zarnowitz (1969) approach which requires the estimation of the coefficients of a regression of the target on a constant and a time series of forecasts.

$$\sigma_t = \alpha + \beta h_t + e_t$$

The null hypothesis is that of $\alpha = 0$ and $\beta = 1$. The MZ regression allows to evaluate two different aspects of the volatility forecast. The MZ regression allows to test the presence of systematic over or under-predictions that is whether the forecast is biased, by testing the joint hypothesis. Being the R squared of an indicator of the correlation between the realization and the forecast, it can be used as evaluation criterion of the accuracy of the forecast.

-Table 7-

To ascertain the accuracy of the forecast we can take as example the case of the UK. We can confirm that the combination of the three variables perform the best. The α equals to zero, the β is 0.997 very close to 1 and the r squared is over 87%.

-Table 8-

In the case of Portugal the RSVOL perform better over the benchmark and the other variables. The R^2 is over 70%, the α is close to zero and the β very close to 1 so as the MZ test dictates the forecast is accurate,

4. Concluding remarks

In the preceding analysis we have presented the use of bridge models in order to nowcast the GDP growth rate of Greece, Cyprus, Ireland, Germany, Portugal, Spain and UK. We found that it is possible to get reasonably good estimates of current quarterly GDP growth in anticipation of the official release. Our results showed that changing the BM's equations by including newly available monthly information provides generally more precise forecasts and is preferable to maintaining the same equation over the exercise's horizon.

Comparing the BMs with DFMs and the MIDAS approach is in our research agenda. Moreover, it would be very interesting to expand the number of explanatory variables to include other economic activity indicators, experiment with different lags of the explanatory variables and, more importantly, with the timing of the monthly releases before the GDP quarterly release. Our goal is to produce forecasts of the GDP and examine their real time performance.

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Table 1

Data series used in our analysis

Data series	Full-sample period	Data collection period/ reporting frequency	Number of observations with reporting lag of 1 month or quarter	Number of observations with reporting lag of 2 months
GDP	1Q 2001-4Q 2013	Quarterly	52	
Industrial production index	Mar 2001-Dec 2013	Monhtly	106	
Volume of Retail Sales	Mar 2001-Dec 2013	Monhtly		106
Total Turnover of Retail Sales	Mar 2001-Dec 2013	Monhtly		106

Source: ELSTAT for Greece and Eurostat for Cyprus, Ireland, Germany, Portugal, Spain and UK

Table 2

Summary of statistics for Greece

	Average	Std. Dev.	Min	Max	Skewness	Kurtosis	ACF(1)	ACF(2)	Correlation with GDP
Real GDP Growth	0,0024	0,049	-0,0894	0,0754	-0,4102	1,7875	0,9199	0,8664	1
IPI (0)	-0,0249	0,047	-0,1312	0,0748	-0,3578	2,5577	0,426	0,4724	0,5804
IPI (-1)	-0,0273	0,0401	-0,1183	0,0513	-0,2611	2,3279	0,5382	0,4874	0,6085
IPI (-2)	-0,0256	0,0482	-0,1403	0,0654	-0,4734	2,6163	0,4734	0,5342	0,6301
RSTOT (0)	0,0174	0,0941	-0,1791	0,1813	-0,4599	1,9496	0,8105	0,6013	0,8388
RSTOT (-1)	0,0212	0,0838	-0,1627	0,1317	-0,5707	1,9758	0,8189	0,7412	0,8663
RSTOT (-2)	0,0165	0,0931	-0,1702	0,158	-0,5672	2,0213	0,736	0,6443	0,7829
RSVOL (0)	-0,0082	0,0848	-0,19	0,1359	-0,5483	2,2405	0,8335	0,6148	0,8617
RSVOL (-1)	-0,0062	0,0773	-0,1635	0,0952	-0,5385	1,841	0,7867	0,7361	0,8789
RSVOL (-2)	-0,0115	0,0854	-0,1755	0,123	-0,4754	1,8701	0,7075	0,6337	0,8175

The variables (0),(-1)(-2) refer to the growth rates of the current month, the previous and two months back, respectively.

Table 3	
ME MAE MSE RMSE for the forecast for the period	200803-201304

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Model	AR (1)	AR(AIC)	RSVOL	RSTOT	IPI	RSVOL&IPI	RSTOT&IPI	ALL 3
ME	-0,008	-0,001	-0,003	-0,004	-0,010	-0,005	-0,007	-0,007
MAE	0,018	0,019	0,017	0,018	0,018	0,016	0,017	0,015
MSE	0,001	0,001	0,000	0,000	0,001	0,000	0,000	0,000
RMSE	0,023	0,023	0,021	0,022	0,023	0,020	0,022	0,019
Ratio 1	1,000	1,006	1,085	1,030	0,993	1,117	1,049	1,192
Ratio 2	0,994	1,000	1,079	1,024	0,987	1,111	1,044	1,185

Ratio1 and Ratio2 are computed as the ratios between each RMSE with that obtained from the AR(1) and AR(AIC) models, respectively.

Table 4

Diebold-Mariano tests of the forecast accuracies of different methods with the benchmark AR(1) and AR(AIC) for Greece

Model	RSVOL	RSTOT	IPI	RSVOL&IPI	RSTOT&IPI	ALL 3
Benchmark the AR(1)	1,11	0,56	-0,07	1,11	0,43	1,36
Benchmark the AR(AIC)	0,76	0,26	-0,08	0,74	0,27	1,23

Table 5

ME,MAE,MSE,RMSE for the forecasts for the period 2008Q3-2013Q4 for Spain,Germany and UK

	AR(1)	AR(AIC)	RSVOL	RSTOT	IPI	RSVOL&IPI	RSTOT&IPI	ALL 3
				Spain				
ME	0,0045	-0,0014	-0,0005	-0,0005	-0,0012	-0,0007	-0,0005	-0,0005
MAE	0,0107	0,0068	0,0061	0,0064	0,0073	0,0070	0,0070	0,0068
MSE	0,0002	0,0001	0,0001	0,0001	0,0001	0,0001	0,0001	0,0001
RMSE	0,0139	0,0094	0,0085	0,0088	0,0090	0,0086	0,0087	0,0086
Ratio 1	1,0000	1,4795	1,6301	1,5686	1,5461	1,6121	1,6032	1,6047
Ratio 2	0,6759	1,0000	1,1018	1,0602	1,0450	1,0896	1,0836	1,0846
				Germany				
ME	0,0004	-0,0016	-0,0015	0,0005	0,0025	0,0024	0,0047	0,0033
MAE	0,0153	0,0156	0,0156	0,0153	0,0145	0,0136	0,0142	0,0140
MSE	0,0004	0,0004	0,0004	0,0004	0,0003	0,0003	0,0003	0,0003
RMSE	0,0203	0,0200	0,0201	0,0196	0,0185	0,0166	0,0169	0,0168
Ratio 1	1,0000	1,0176	1,0096	1,0382	1,1010	1,2253	1,2041	1,2122
Ratio 2	0,9827	1,0000	0,9922	1,0203	1,0820	1,2041	1,1834	1,1913
				UK				
ME	0,0018	-0,0027	-0,0019	-0,0025	-0,0014	-0,0009	-0,0013	0,0003
MAE	0,0117	0,0090	0,0088	0,0090	0,0084	0,0083	0,0084	0,0079
MSE	0,0002	0,0001	0,0001	0,0001	0,0001	0,0001	0,0001	0,0001
RMSE	0,0158	0,0111	0,0109	0,0113	0,0108	0,0105	0,0107	0,0098
Ratio 1	1,0000	1,4204	1,4411	1,3975	1,4623	1,5054	1,4755	1,6156
Ratio 2	0,7040	1,0000	1,0146	0,9839	1,0295	1,0598	1,0388	1,1374
Table 6								

ME,MAE,MSE,RMSE for the forecasts for the period 2008Q3-2013Q4 for Portugal, Cyprus and Ireland

	AR(1)	AR(AIC)	RSVOL	IPI	ALL 2
		Р	ortugal		
ME	-0,0022	-0,0050	-0,0025	-0,0053	-0,0020
MAE	0,0109	0,0101	0,0098	0,0108	0,0105
MSE	0,0002	0,0001	0,0001	0,0002	0,0001
RMSE	0,0139	0,0122	0,0120	0,0125	0,0122
Ratio 1	1,0000	1,1369	1,1573	1,1095	1,1367
Ratio 2	0,8796	1,0000	1,0179	0,9759	0,9998
		(Cyprus		
ME	-0,0060	-0,0049	-0,0065	-0,0060	-0,0059
MAE	0,0110	0,0101	0,0093	0,0091	0,0093
MSE	0,0002	0,0001	0,0001	0,0001	0,0001
RMSE	0,0135	0,0122	0,0111	0,0112	0,0111
Ratio 1	1,0000	1,1078	1,2186	1,2075	1,2131
Ratio 2	0,9027	1,0000	1,1000	1,0900	1,0950
		1	reland		
ME	-0,0108	-0,0018	-0,0048	-0,0060	-0,0070
MAE	0,0235	0,0287	0,0292	0,0302	0,0307
MSE	0,0011	0,0016	0,0014	0,0015	0,0016
RMSE	0,0324	0,0395	0,0376	0,0392	0,0401
Ratio 1	1,0000	0,8216	0,8628	0,8280	0,8079
Ratio 2	1,2171	1,0000	1,0501	1,0078	0,9833

Table 7				
Mincer-Zarnowitz	test for	Spain.C	Jermany.	.UK

		······································				
		Alpha	p-value	Beta	p-value	R2
	AR(1)	0,001	0,385	0,664	0,000	0,899
	AR(AIC)	-0,002	0,154	0,772	0,000	0,911
P.	RSVOL	-0,001	0,302	0,788	0,000	0,929
pai	RSTOT	-0,001	0,323	0,785	0,000	0,919
	IPI	-0,002	0,215	0,796	0,000	0,908
	RSVOL&IPI	-0,002	0,270	0,789	0,000	0,924
	RSTOT&IPI	-0,001	0,348	0,795	0,001	0,917
	ALL 3	-0,001	0,367	0,774	0,000	0,936
		Alpha	p-value	Beta	p-value	R2
	AR(1)	0,001	0,768	0,869	0,362	0,633
~	AR(AIC)	0,000	986,000	0,831	0,196	0,661
any	RSVOL	0,000	0,942	0,801	0,111	0,669
rm	RSTOT	0,002	0,699	0,836	0,200	0,672
Ge	IPI	0,003	0,336	0,808	0,079	0,732
	RSVOL&IPI	0,003	0,329	0,831	0,830	0,784
	RSTOT&IPI	0,005	0,138	0,861	0,172	0,777
	ALL 3	0,004	0,246	0,841	0,111	0,778
		Alpha	p-value	Beta	p-value	R2
	AR(1)	0,001	0,775	0,732	0,004	0,778
	AR(AIC)	-0,003	0,251	0,924	0,363	0,852
	RSVOL	-0,002	0,412	0,913	0,293	0,853
UK	RSTOT	-0,001	0,293	0,918	0,339	0,846
	IPI	-0,001	0,528	0,963	0,676	0,849
	RSVOL&IPI	-0,001	0,686	0,953	0,575	0,857
	RSTOT&IPI	-0,001	0,578	0,977	0,794	0,851
	ALL 3	0,000	0,897	0,997	0,974	0,873

Table 8

Mincer-Z	Mincer-Zarnowitz test for Portugal, Cyprus and Ireland									
		Alpha	p-value	Beta	p-value	R2				
al	AR(1)	-0,004	0,257	0,839	0,2964	0,586				
gug	AR(AIC)	-0,006	0,026	0,875	0,286	0,727				
ort	RSVOL	-0,004	0,135	0,827	0,138	0,712				
H	IPI	-0,006	0,031	0,913	0,492	0,708				
	ALL 2	-0,003	0,22	0,843	0,211	0,687				
		Alpha	p-value	Beta	p-value	R2				
S	AR(1)	-0,006	0,023	0,916	0,314	0,852				
pru	AR(AIC)	-0,005	0,013	0,822	0,004	0,91				
Cyl	RSVOL	-0,007	0,001	0,89	0,049	0,929				
	IPI	-0,006	0,004	0,896	0,083	0,917				
	ALL 2	-0,006	0,004	0,908	0,134	0,915				
		Alpha	p-value	Beta	p-value	R2				
q	AR(1)	-0,011	0,084	0,628	0,056	0,347				
an	AR(AIC)	-0,007	0,267	0,419	0	0,312				
[re]	RSVOL	-0,008	0,231	0,428	0,006	0,192				
	IPI	-0,009	0,193	0,405	0,002	0,206				
	ALL 2	-0,009	0,178	0,378	0,002	0,172				

Appendix B.



Graphs of each time series in comparison with the Real GDP Growth.









